



## MACROECONOMIC POLICY RESPONSES TO COVID-19

MEASURING FINANCIAL RESTRICTIONS OF  
BRAZILIAN PRIVATE FIRMS WITH MICRODATA:  
DID CREDIT POLICIES OF BANCO CENTRAL  
DO BRASIL IN COVID-19 PANDEMIC AFFECT  
INVESTMENT DEMAND?

By Fernando Oliveira

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# Measuring Financial Restrictions of Brazilian Private Firms with Microdata: Did Credit Policies of Banco Central do Brasil in Covid-19 Pandemic Affect Investment Demand?\*

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## Abstract

Our contributions in this paper are twofold. In the first place, we build original financial restriction measures of 8,071 private firms in Brazil. We use these measures to have an idea of their level of financial restrictions from 2012 to 2020. To build these measures, we use microdata of 3,469,135 loan contracts written between these firms and financial institutions in this period. In the second place, we estimate, using our financial restriction measures, investment demand of these firms and verify if credit policies of the Banco Central do Brasil during the Covid-19 pandemic had any positive effect in mitigating their credit restrictions. Our results show that our financial restriction measures explain well the access of Brazilian private firms to credit for investment as well as indicate that their investment is negatively related to financial restrictions in Brazil. Furthermore, they also show that credit policies of Banco Central do Brasil had positive effect on working capital loans of private firms but did not have any effect on their investment in the Covid-19 pandemic period.

**JEL Codes:** G13, G32, G38.

**Keywords:** financial restrictions measures, investment, private firms, Covid-19 pandemic, Banco Central do Brasil.

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\* The opinions in this paper are those of the author and not necessarily reflect the official opinions of Banco Central do Brasil on the issue.

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# 1 Introduction

Financial restrictions are a key and widespread concern for firms, thwarting their ability to carry out optimal investment policies and growth trajectories. They are very important both in academia and for policy makers. In the case of the former, they are very relevant both in theoretical and empirical work related to Macroeconomics, Corporate Finance, Monetary Economics, Game Theory and Economics of Contracts. In the case of latter, they are relevant to have an idea of the stance of the credit market in the economy.

The existing theoretical literature on the effects of financing constrains on firms' performance is very extensive. However, empirical contributions are rather scarce and more recent. In addition, despite its enormous importance, measuring financial restrictions is still subject to much discussion, because most empirical studies have not only to deal with a set of measurement and conceptual issues, but also rely on tenuous relationships between firms and banks to identify the presence and severity of financial constraints.

The very definition of financial restrictions is not clear-cut. One can address several aspects of the relationship between firms and banks to extract measures of financial restrictions.<sup>1</sup> One common feature of all financial restriction measures, however, is the one that differentiates them from financial distress measures. A firm can be healthy and efficient in economic and financial terms, therefore not in distress, and still be credit restricted for some reasons.

In this paper, we will use a definition of financial restrictions related to the capacity a firm has to obtain credit with banks to implement positive present value investments or projects.<sup>2</sup> So, just to exemplify and make our definition clearer, let us suppose a certain firm has a positive present value investment and does not have all the monetary resources necessary to undertake this investment. The firm asks for loans in several banks for this purpose and all banks deny it. In accordance with our definition, this firm is financially constrained.

Of course, an empirical researcher wanting to verify if the firm is financially restricted or not would have enormous difficulty doing so, looking only at the loan data of this firm. To have a better grasp of this, she should have to ask the firm if it has a positive present value investment that it wants to implement and needs credit from banks for this and ask each bank that denied the credit for this purpose why it did so. This makes financial restrictions in empirical work non-observable and hence very difficult to measure. No wonder, empirical researchers have been struggling for a very long time and devoting a lot of hard work in trying to measure financial restrictions of firms in appropriate

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<sup>1</sup>Kaplan & Zingales (1997) use a broader definition by stating that financial constraints are present whenever there is a wedge between the costs of obtaining internal and external funds. However, the problem with such definition is that it almost covers every firm.

<sup>2</sup>This definition is a very common one in the empirical literature that measures financial restrictions (see Fazzari et al. (1988) for a discussion).

manners.

Our main objective in this paper is to create financial restriction measures based on microdata related to bank loan contracts of Brazilian private firms. To understand how our measures of financial restrictions contribute to the empirical literature, it is important to describe very briefly the current state of the art of the empirical literature on this issue. As [Carreira & Silva \(2012\)](#) point out, there are three types of financial restriction measures: indirect measures, direct measures and indexes.

Indirect measures look at the sensitivity of investment in relation to cash flow. The fundamental idea is that if sensitivity is high, then there is an indication that the firm may be financially restricted. To test for this possibility, firms are separated into two groups—more likely to be financially restricted and less likely to be financially restricted—based on ex-ante classification related to balance sheet characteristics of firms.

The seminal empirical paper of indirect measures is by [Fazzari et al. \(1988\)](#). The authors use firm's dividend policy to classify them in financially restricted or not. Firms that pay more dividends would be less likely, while firms that pay fewer dividends would be more likely to be financially constrained. The intuition behind this classification is that firms that pay fewer dividends use more internal resources to invest, since they have credit restrictions.<sup>3</sup>

There are many other ex-ante classification possibilities of financial restrictions based on balance sheet information of firms. One that is widely used and has the advantage of being exogenous is size (see [de Oliveira \(2019\)](#) for Brazil's data and [Campello et al. \(2010\)](#)).<sup>4</sup>

Instead of looking at investment cash flow sensitivities, [Almeida et al. \(2004\)](#) look at the cash policy of firms and analyze the cash sensitivity to cash flow. The main idea of [Almeida et al.](#) is that if a firm hoards cash, then it is probably because it wants to have enough liquidity to undertake its investments without asking banks for credit.<sup>5</sup>

Indirect measures have some drawbacks. The main concern is that all of them are associated with the average Q of Tobin. This firm multiple is used to identify investment opportunities. The average Q of Tobin is a proxy for the marginal q of [Tobin \(1969\)](#), which the neoclassical theory of investment considers being the correct measure to identify the set of investment opportunities. The problem is the average Q may be a bad proxy for the marginal q, as [Erickson & Whited \(2000\)](#) show. It may also be the case, as [Cleary et al.](#)

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<sup>3</sup>[Kaplan & Zingales \(1997\)](#) question this classification and show that many firms in the sample they used paid very little dividend but had no indication of being financially constrained. The authors cite Hewlett-Packard as an example of a firm of this sort.

<sup>4</sup>The most common way to measure the size is by total assets, although some authors also use the number of employees or age. In general, however, age and number of employees are highly correlated with size.

<sup>5</sup>There are other balance sheets or financial characteristics used in the indirect measures, such as credit ratings, or firms' capacity to give collaterals for loans, among others. See [Carreira & Silva \(2012\)](#) for more details.

(2007) stress that cash flow may also contain information on investment opportunities when there is high uncertainty about a firm's investment projects.

As it is impossible to measure marginal  $q$  correctly, this makes tests of investment cash flow sensitivity based on indirect measures imprecise in statistical terms. Another problem with the average  $Q$  of Tobin is the fact that it only exists for listed firms. However, listed firms are the ones that would normally have more access to credit. In contrast to listed firms, private firms are the ones that should be more prone to depend on banks for investment.

Direct measures of financial restrictions, different from indirect measures, do not use the average  $Q$  of Tobin. They are built on surveys and reports of firms. They are firm-specific, time-varying, and researchers can use them as dependent or independent variables in their studies. For instance, in the case of reports, [Kaplan & Zingales \(1997\)](#) read them, searching for expressions that are symptomatic of the presence of financial constraints.

Another possibility to create direct measures of financial restrictions is through surveys.<sup>6</sup> In surveys, firms are asked whether they are financially restricted or not and this can be done by a single question or combination of different questions, related to their cost of external funds, credit denials, and availability of external funds. The main advantage of surveys is that firms are the best-informed agents with respect to the quality of their projects. One should expect that investment opportunities are already considered in the firm's responses. It is also possible to measure financial restrictions for small and young firms that do not publish their balance sheets, which is not possible in the case of indirect measures. To complement survey information, one can also use quantitative information as well.

Like indirect measures, direct measures also have some downsides. The subjective nature of self-assessed variables means that potential biases resulting from management perceptions may exist. In addition, information is expensive to collect, somewhat scarce, with an insufficient level of detail. It is also important to complement surveys with information coming from financial institutions, which is seldom available.

The third type of measure of financial restrictions is an index. This is a combination of direct and indirect measures. [Kaplan & Zingales \(1997\)](#) and [Whited & Wu \(2006\)](#) are some of the most important and cited indexes in the literature. They are built based on qualitative and quantitative information and share the advantages and disadvantages of direct and indirect measures.

We think that our paper contributes in an important way to the empirical literature, because not only our measures of financial restrictions are different, but also, more accurate than the ones that exist in the literature so far and that we describe briefly above.

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<sup>6</sup>See, for example, [Campello et al. \(2010\)](#), [Beck et al. \(2006\)](#) and the [European Central Bank Access to Finance Survey](#).

The reason we think like this is that we construct them by the analyzing credit information of private firms directly from their loan contracts with financial institutions. We use information from 3,469,135 loan contracts of 8,010 private firms written with banks from 2012 to 2020 and registered at the System of Credit Registry (hereafter SCR) of Banco Central do Brasil (hereafter BCB).

The SCR allows us to have a better comprehension of the access of credit of firms for investment. So, for example, we know that the firm is in liquidation or in bankruptcy in which case it will be very unlikely to obtain credit for investment. SCR also informs the motive or type of the loan contract. There is information that the firm obtained credit for investment or for project financing in which case we also can deduce that the firm is not credit-constrained for investment. There is also very detailed information on the maturity and interest rates of loans, among other very relevant credit information.

We classify a private firm every year of our sample in five different categories depending on its likelihood to obtain credit for investment: very likely to be financially restricted (liquidation or restructuring), likely to be financially restricted, not capable of identifying, likely not to be financially restricted and very likely not to be financially restricted (in the case the firm obtained credit for investment or project financing).

Our core measures of credit restriction classify private firms in a certain year as likely to be financially restricted or not, depending on how the average interest rate and average maturity of their contracts fit in the cross-section distributions of interest rates and maturities of all loan contracts of firms written in this year with financial institutions.

We build also other measures that complement our core measures with information of private firms related to: the number of financial institutions they have a relationship with; if they have loans in delinquency over 90 days; the quality of their loans portfolios; if they have outstanding foreign exchange derivatives contracts<sup>7</sup>; how much collateral they offer in their loan contracts; and, finally, some balance sheet characteristics such as coverage ratio, fixed assets and size (total assets).

We also think that we contribute to the literature because we study only private firms, which is also not common in the literature. Most papers that measure financial restrictions look only at listed firms, which should be much less likely to be credit constrained for investment.

Our financially restricted measures have most of the desired properties [Carreira & Silva \(2010a\)](#) mention, such as being simple, objective, firm-specific and time-varying. In addition, given our financial restricted measures we may understand better investment cash flow sensitivity in Brazil.

Our results show that our financial restriction measures explain well access of private

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<sup>7</sup>We use information of foreign exchange derivatives contracts, because we ponder that a firm that has access to these derivatives, either through a Mercantile Exchange or through over-the-counter contracts, is less likely to be financially constrained.

to credit for investment as well as indicate that their investment is negatively related to financial restrictions in Brazil. Furthermore, our results also show that credit policies of BCB had a positive effect on working capital loans but did not have any effect on investment of private firms during the Covid-19 pandemic period.

The remainder of the paper is the following. Section 2 describes in more details the empirical literature on financial constraint measures. Section 3 describes the data. Section 4 presents our empirical identification strategy. Section 5 shows the results of the main and robustness empirical analyses. Section 6 concludes.

## 2 Literature Review of Financial Constraints Measures of Firms

The theoretical literature on measures of the relationship between financial constraints of firms and investment is based on the relaxation of the hypothesis of perfect markets of Modigliani & Miller (1958)'s theorem. Modigliani & Miller demonstrate that, in perfect capital markets, external finance is a perfect substitute for internal finance, thus the financial structure of firms and financial policy are irrelevant for their investment decisions.

Stiglitz & Weiss (1981) and Myers & Majluf (1984) document well the consequences of capital market imperfections due to agency problems. The fundamental idea is based on the existence of moral hazard and adverse selection problems that either set the price of credit above optimal levels or rationalize (in some circumstances by complete) credit. This inefficiency sets a wedge between internal and external forms of firms' financing. As a result, firms' investment decisions will not be optimal, and they will not be able to fulfill their growth and investment optimal targets.

The problems with asymmetric information in capital markets can be more severe for small and young firms. This will happen because there is still not much information on these firms available to the most potential lender. Potential lenders are not able to observe the quality of the risk or do not have control over the firm's investment. Under these conditions, smaller and younger firms are expected to be more credit constrained, as shown by Petersen & Rajan (1994, 1995).

Empirical work related to the measurement of financial restrictions and its relation to investment can be decomposed into three distinct strands, as we explained very briefly in the Introduction: indirect measures, direct measures and indexes. Indirect measures are related to the study of the sensitivity of investment to cash flow; direct measures are based on surveys and off-balance sheet reports of firms, while indexes are based on the combination of direct and indirect approaches.

To measure the sensitivity of investment to cash flow, indirect measures use regressions

based on investment as a dependent variable and average Q of Tobin and cash flow as explanatory variables, besides separating the sample of firms in credit-constrained and non-constrained firms. The null hypotheses of these tests are that cash flow sensitivity is going to be higher for constrained firms than for unconstrained firms.

To separate, ex-ante, firms in financially restricted or not, the literature uses several balance sheet or financial characteristics such as: assets, age, number of employees, credit ratings, number of bank relationships, among some others.<sup>8</sup>

Average Q of Tobin is a proxy for marginal q of Tobin (1969), which measures the increase in the present value of a firm's profits resulting from a marginal increase in the firm's capital stock and is non-observable.<sup>9</sup> There are several ways to define average Q. One that is very common in the literature is the ratio between the market value of the firm and its cost of capital replacement. As this cost of capital replacement is not observable, total assets or fixed assets are often used as substitutes in the literature. A high level of Q thus would be an indication of the presence of investment opportunities. It is argued that Q (or marginal q being more precise) summarizes all future information that is relevant for a firm when deciding to invest.

A relevant setback of the average Q of Tobin is the fact that it only exists for listed firms. An interesting alternative to the estimation of investment without Q of Tobin is developed by Gala et al. (2020). The authors estimate investment demand without information about market values and hence the average Q of Tobin, but under general assumptions about technology and markets.

The seminal work on indirectly measured financial restrictions is by Fazzari et al. (1988). The authors investigate the impact of cash flow sensitivities on investment by classifying firms as credit-constrained or not according to their dividend policy. The reason for this classification rests on the argument that firms that pay low dividends, since their needs for resources for investment exceed their internal cash flow and they are credit constrained. They show using a sample consisting of 422 USA firms from 1970 to 1984 that the coefficient of cash flow for the low-dividend group is higher and statistically different than the coefficient for the high-dividend group. This suggests that low-dividend firms invest more of their extra cash flow than high-dividend firms.

On the other hand, Kaplan & Zingales (1997) argue that cash flow is not a good measure of the existence of financing constraints and Fazzari et al. (1988) a priori classification of firms is flawed. They instead classify firms according to information obtained from company annual reports and find evidence that constrained firms are the less sensitive to cash flow. This argument is also supported by Kadapakkam et al. (1998) and Cleary (1999). Recently, Dasgupta & Sengupta (2007), for Japan, find that the response of investment to cash flow shocks is non-monotonic, supporting Kaplan & Zingales (1997)

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<sup>8</sup>See Carreira & Silva (2010b)

<sup>9</sup>Hayashi (1982) demonstrates that in perfect capital markets average Q is equal to marginal q.



and [Cleary \(1999\)](#).

Alternatively, analyzing firms' demand for cash, [Almeida et al. \(2004\)](#) claim that the level of financial constraints can be measured by the sensitivity of cash stock to cash flow. The rationale behind is that, while constrained firms need to save cash out of cash flows in order to take advantage of future investment opportunities, unconstrained firms do not, as they are able to resort to external finance. Meanwhile, firms that hold cash incur in opportunity costs associated with present investment opportunities. As a result, only constrained firms will need to optimize their cash stocks over time in order to maximize their profits and hedge future socks by holding cash. Therefore, one can expect that estimates on the sensitivity of cash stocks to cash flow.

Another facet of the literature points out that cash flows might contain information about firms' investment opportunities, meaning that Q should be corrected, as [Alti \(2003\)](#) and [Bhagat et al. \(2005\)](#) point out. [Alti](#) finds that even after Q correction, every firm in his sample shows sensitivity to cash flow. In addition, [Bhagat et al.](#) find evidence that financially distressed firms exhibit positive investment-cash flow sensitivities if they operate at a profit, and low sensitivity if they operate at a loss.

Close bank relationships facilitate the contact between firms and banks, reducing information asymmetries, which means lower financing constraints for firms (if such relationships are stable). As [Diamond \(1991\)](#) argues, the risk associated with any loan is not neutral with respect to the duration of the relationship. As a result, one can expect differences in financial constraints between market-oriented economies (such as the USA and the UK) and bank-oriented ones (Germany for example).

An interesting empirical paper that studies firms' bank relationships is [Karaivanov & Townsend \(2014\)](#). The authors examine whether financial constraints affect firms' investment decisions by comparing a group of unbanked firms to firms that rely on bank financing. Specifically, they combine data from the Spanish Mercantile Registry and the Bank of Spain Credit Registry (CIR) to classify firms according to their number of banking relations: one, several, or none. They show that financial constraints are negatively related to the number of bank relationships firms have.

In the case of direct measures, one can read the annual reports of firms and look for words or expressions of words that give some hint of the financial difficulties a firm is facing such as [Kaplan & Zingales \(1997\)](#) did. Otherwise, one can prepare surveys for firms to answer one or more questions related to their cost of external funds, credit denials, and availability of external funds, as in, for example, [Campello et al. \(2010\)](#), [Beck et al. \(2006\)](#), [European Central Bank's survey on the access to finance of enterprises \(SAFE\)](#) and [Ferrando & Mulier \(2015\)](#).

[Campello et al. \(2010\)](#) surveyed 1,050 chief financial officers (CFOs) in 39 countries in North America, Europe, and Asia in December 2008. They contrast the actions of firms that are financially constrained with those that are less constrained. They develop

a survey-based measure of financial constraint and then study whether this constraint measure identifies meaningful cross-sectional variation in corporate behavior during the crisis.

[Beck et al. \(2006\)](#) examined whether financial development boosts the growth of small firms more than large firms and hence provides information on the mechanisms through which financial development fosters aggregate economic.

The [European Central Bank's](#) survey on the access to finance of enterprises (SAFE) provides information on the latest developments in the financial situation of enterprises, and documents trends in the need for and availability of external financing. The survey results are broken down by firm size, branch of economic activity, country, firm age, financial autonomy and ownership. The survey is conducted twice a year.

[Ferrando & Mulier \(2015\)](#) draw on the SAFE survey of 11,886 firms in the euro area to investigate the role of firm characteristics with respect to the experience of facing financial restrictions from 2009 to 2011. Their methodology is based on nearest neighbor matching of the firms in their sample with balance sheet information of 2.3 million firms in the euro area. They are capable of distinguishing perceived from actual financial obstacles and they show that more profitable firms are less likely to face actual financing constraints.

The third strand of the literature is on indexes. They combine indirect and direct measures, and thus have the advantages and disadvantages of them. They have quantitative as well as qualitative information. Some of the most important indexes in the literature are [Kaplan & Zingales \(1997\)](#) and [Whited & Wu \(2006\)](#). The Kaplan-Zingales index is a relative measurement of reliance on external financing. Companies with high index are more likely to experience difficulties when financial conditions tighten since they may have difficulty financing their ongoing operations. The index is based on a five-factor model of the following variables: cash flow, Q of Tobin, Total Debt, dividends and Cash.

[Whited & Wu \(2006\)](#) derived their index from a generalized method of moments estimation of an investment Euler equation. The Lagrange multiplier of the external financing constraint is the shadow cost of external financing. It is a function of five factors: the ratio of total debt to total assets, an indicator if the firm pays dividends or not, the growth of firm's sales, the firm's sector growth and cash divided by total assets.

The existence of financing constraints appears to be particularly severe for firms that decide to invest in RD because of the risks associated with the investment. As argued before, credit markets will no longer be efficient, generating a wedge between internal and external financing faced by firms as well as a financing hierarchy. For example, [Hall \(1992\)](#) and [Himmelberg & Petersen \(1994\)](#), find support for the hypothesis that RD investment is financially constrained for small firms. [Hall et al. \(1999\)](#) in a comparative study of French, Japanese, and American firms, also sustain these findings.

The financial constraints faced by firms can, obviously, have important effects on

the firm’s ability to stay in the market. For example, [Musso & Schiavo \(2008\)](#) find that, for French manufacturing firms over the period 1996-2004, the greater the financial constraints firms face, the higher the probability that they do not survive and then exit the market.

### 3 Data

We have two sources of data. The balance sheet and financial information of private firms (hereafter firms) come from Valorpro.<sup>10</sup> Valorpro has the advantage of having information mostly on medium to small size firms, which are our main interest in this paper. The information on firm’s loan contracts comes from SCR.

Our database has unbalanced balance sheet and financial information of 8,010 firms from 2012 to 2020. There are 5,503 joint stock and 2,568 limited liability firms. We classify these firms into 5 sectors, following the classification scheme of Valorpro: Agriculture, Commerce, Energy, Industry and Services. Most firms come from the services sector (4,111) (see [Table 1](#)).

Table 1: Firms and Sectors

Sectors	Joint Stock	Limited Liability	Total
Agriculture	179	16	195
Commerce	573	217	790
Energy	778	218	996
Industry	1,358	621	1,979
Services	2,615	1,496	4,111
Total	5,503	2,568	8,071

*Notes:* Our database of firms has unbalanced balance sheet and financial information of 5,503 joint stock private firms and 2,568 limited liabilities private firms from 2012 to 2020. We classify these firms in 5 sectors, following the classification of Valorpro: Agriculture, Commerce, Energy, Industry and Services.

[Table 2](#) Panel A displays the descriptive statistics (mean and standard deviation) of balance sheet and financial information firms in our sample. The largest ones (measured by the natural logarithm of total assets) come from the energy sector followed by the industry sector. The highest average investment (capex) occurs in the services sector (10%). The most profitable sector, measured by the average quotient between ebitda and total assets, is agriculture sector (1.46), and the one with the highest leverage corresponds to the energy sector (0.95). In terms of annual growth of operational revenues, on average, the services sector has the highest average value of 3.93%.

[Table 2](#) Panel B shows the number of foreign exchange derivatives contracts of our database firms classified by type of derivative, long or short positions and sectors. Forward

<sup>10</sup>Valorpro is a database of balance sheet and financial information of firms. It is a proprietary database of Brazilian economic journal Valor Econômico

contracts are predominant, followed by option contracts. Most firms are long in the foreign exchange rate.

Table 2: Descriptive Statistics

Panel A: Mean and Standard Deviation Private Firms

	Agriculture	Commerce	Energy	Industry	Services
logassets	4.1977	3.8509	4.4828	4.1586	4.1088
	0.8640	0.8897	0.7940	0.9078	0.8980
capex/assets	0.0581	0.0427	0.0617	0.0256	0.1009
	0.2075	0.0961	0.8859	0.1208	3.2642
ebitda/assets	1.4633	1.1858	1.0119	0.8821	1.3840
	8.4061	7.4990	9.9141	27.6390	32.6111
rev_oper/assets	0.0439	0.0473	0.0727	0.0437	0.0958
	0.1296	0.1580	0.4915	0.2330	9.9161
var_oper_rev	-0.5465	0.2640	-3.4688	0.5434	3.9330
	32.7163	12.9743	224.4119	110.5750	312.5313
leverage	0.8257	0.7470	0.9571	0.9389	0.6182
	0.5624	0.6083	0.5617	0.5119	0.5915

Panel B: Number and Types of Foreign Exchange Derivatives Contracts

	Forward		Options		Swap		Future	
	Long	Short	Long	Short	Long	Short	Long	Short
Agriculture	5,165	6,857	86	102	241	92	0	0
Commerce	41,106	43,632	773	609	2,941	426	7,739	7,684
Energy	6,863	9,732	512	523	1,550	300	0	0
Industry	19,320	21,013	11,598	10,534	2,599	1,276	0	0
Services	4,724	1,822	173	98	1,529	127	0	0
Total	77,178	83,056	13,142	11,866	8,860	2,221	7,739	7,684

*Notes:* Our database of firms has unbalanced balance sheet and financial information of 5,503 joint stock private firms and 2,568 limited liabilities private firms from 2012 to 2020. We classify these firms in 5 sectors, following the classification of Valorpro: Agriculture, Commerce, Energy, Industry and Services. Panel A shows mean and standard deviation (second number below) of balance sheet characteristics separated by sectors of the economy. Panel B shows the number and type of foreign exchange derivatives contracts outstanding from 2012 to 2020 separated by sectors of the economy.

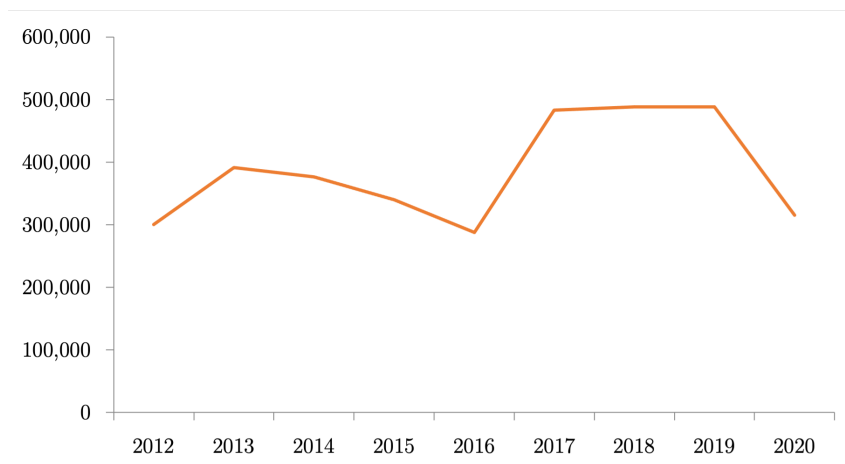
In the case of loan contracts, we use the information of SCR. The SCR is an outstanding database of loan contracts written between individuals, firms and financial institutions. Our interest in this paper is to study the loans of firms.

The SCR has flow, stock, and cadastral information of these loan contracts. In the case of cadastral information, SCR informs if the firm is in liquidation or a restructuring process due to a bankruptcy process. In the case of flow information, one can observe, among other features of the loan contracts, the date the contract was written, the interest rate charged, maturity and motive or type of the loan contract, and the credit risk of the loan. Stock information, for example, relates to if it has more than 90 days delinquency in some loan, the relative importance of bad loan in relation to the total portfolio of loans, the number of bank relationship the firms have, among other characteristics.

In our empirical identification strategy, as we will make clear in the next section, we will use all the information on loan contracts mentioned above. In the case of flow information, we are interested in contracts written between firms and financial institutions from 2012 to 2020.

Figure 1 shows the total number of loan contracts written every year from 2012 to 2020 by firms in our sample registered at SCR. It is interesting to observe there is a sharp decline on the number of contracts between 2014 and 2016. We conjecture that this may have happened because of a decrease in GDP growth, which had to do with deterioration of fiscal side of the economy observed in Brazil during these years.

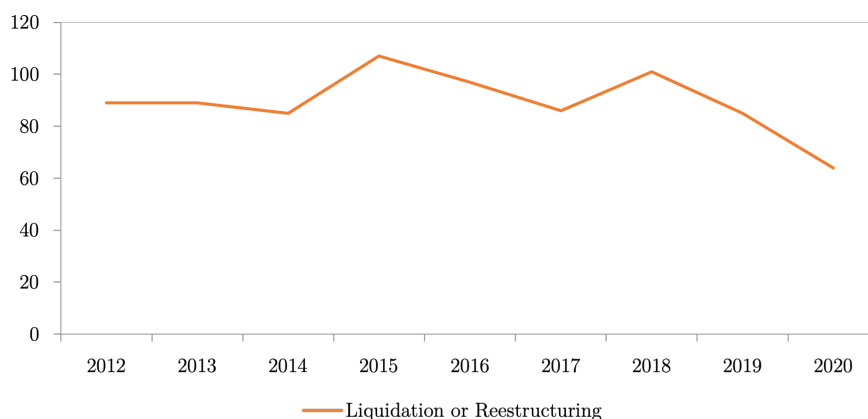
Figure 1: Number of Loan Contracts of Firms



Source: Banco Central do Brasil (BCB) – System of Credit Information (SCR)

The information contained of loan contracts in SCR allows us to distinguish firms that are in liquidation or restructuring due to a bankruptcy process. Figure 2 shows the number of firms in such situations changed little in the period. This number is relatively stable throughout our sample period.

Figure 2: Number of Firms in Liquidation or Restructuring



Source: Banco Central do Brasil (BCB) – System of Credit Information (SCR)

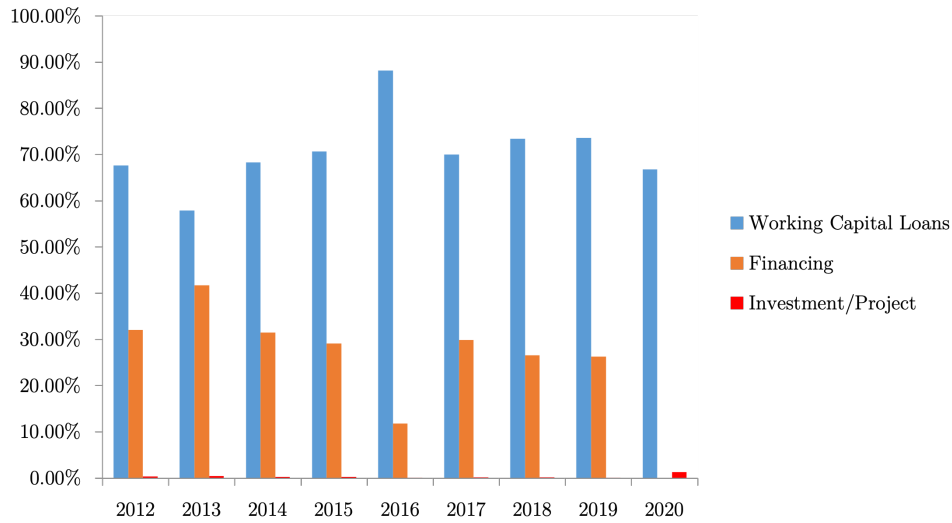
We classify the type or motive of the loan contract into three categories: working capital, financing and investment or project finance. Working capital is all sorts of loans that are not for financing or investment reasons. Examples of these types of loans are bank discount of credit instrument, secure overdraft facilities, long-term working capital, short-term working capital, credit card receivable financing, among many others.

Financing loans, on the other hand, exist for the purchase of goods, whether mobile or not. Normally, and as a rule, interest rates are lower and maturities are higher for financing loans than for those working capital loans, since the financed asset is, normally, given as collateral until the debt is paid off. It is not clear, however, that, in the case firms are acquiring financing loans that they are investing, that is, increasing their capital stock.

In addition, we have information about loans for project financing or investment, which is crucial for our empirical identification strategy, because it indicates firms that are acquiring these loans to increase their capital stock.

Figure 3 presents the percentage of types of loan contracts of the firms in our database from 2012 to 2020. Figure 3 shows that working capital loans exceed by far loans for financing and investment or project financing in all years. It is important to note that almost all loans in 2020—which occurred the Covid-19 pandemic (hereafter pandemic)—were working capital loans. We think that this has to do with the credit policies of BCB implemented during the pandemic to mitigate its effects on the financial restrictions of firms.

Figure 3: Percentages of Types of Loans Contracts of Firms



Source: Banco Central do Brasil (BCB) – System of Credit Information (SCR)

## 4 Empirical Strategy

### 4.1 Identification of Financial Restriction Measures

Our empirical identification strategy of financial restrictions measures is the following. We start by defining five categories of a firm’s relative access to credit for investment. The categories are very likely non-financially restricted, likely to be financially restricted, not enough information to classify, likely to be non-financially restricted, and very likely to be non-financially restricted.

Table 3 presents our financially restricted measures based on the categories above. All our measures are calculated every year, so they can change in our sample period. Our core measures presented in Table 3 Panel A are FR1\_Contracts, FR2\_Contracts, FR3\_Contracts.<sup>11</sup> They are equal to 1 if the firm is in restructuring or liquidation; they are equal to 5 if the firm obtained loans for investment or project financing; they are equal to 3 if we do not have enough information to identify if they are financially restricted or not.

FR1\_Contracts, FR2\_Contracts, and FR3\_Contracts differ in the way they fall into categories 2 or 4. For a certain firm in a certain year, FR1\_Contracts is equal to 4 in two situations. In the first one, a firm acquired financing loans. In the second one, a firm obtained only working capital loans, in which the average interest rate was lower than 30% percentile of its cross-section distribution and average maturity was higher than the 70% percentile of its cross-section distribution.

<sup>11</sup>FR is an acronym for financial restrictions.

FR1\_Contracts is equal to 2 if the firm acquired obtained only working capital loans and the average interest rate of these working capital loans was higher than the 70% percentile of its cross-section distribution, as well as average maturity was lower than the 30% percentile of its cross-section distribution.<sup>12</sup>

FR2\_Contracts and FR3\_contracts differ from FR1\_Contracts when we classify them in categories 2 or 4, because they consider the 80% and 90% percentile respectively of the cross-section distribution of maturities instead of the 70% percentile, and because they use the 20% and 10% percentile of cross section distribution of interest rates respectively instead of 30% percentile.

We build other financial restriction measures that complement our measures above using the information of firms related to the number of delinquency days of loans; the demand for foreign exchange derivatives; the number of bank relationships; information of their portfolio on non-performing loans (bad loans); average collateral provided; and from some balance sheet information. All this information is relevant because it can change the classification of our core measures in category 3 (unable to identify credit restrictions) into categories 2 (likely to be restricted) or 4 (likely to be non-restricted).

Table 3 Panel B describes how these complementary measures allow us to change a firm from category 3 (not enough information) to category 2 (very likely to be constrained) in a certain year: delinquency if we observe that the firm had loans in delinquency over 90 days; a bad portfolio, one in which a firm has more than 70% of its portfolio of loans non-performing; balance sheet information related to total assets, fixed assets in relation to total assets and coverage ratio, where their average is less than the 30% percentile of their respective cross-section distribution; the ratio between average collateral pledged and total assets is less than 30% percentile of its cross-section distribution; finally, the number of bank relationships of firms is less than 30% percentile of its cross-section distribution.

Table 3 Panel C describes other measures that allow us to change a firm from category 3 (not enough information) to category 4 (likely not to be constrained) in a certain year: good portfolio, one in which a firm has less than 30% of its portfolio of loan non-performing; balance sheet information related to total assets, fixed assets in relation to total assets and coverage ratio is higher than the 30% percentile of their respective cross-section distribution; average ratio of collateral pledged to assets is higher than 70% percentile of the its cross section distribution; finally, the number of bank relationships of firms is higher than 30% percentile of its cross section distribution.

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<sup>12</sup>The 70% percentile and 30% percentile are thresholds adopted as a simple back-of-the-envelope rule and are commonly used in many empirical works, see [Moore et al. \(2013\)](#).



Table 3: Financial Restriction Measures and Categories of Financial Restrictions

## Panel A: Core Financial Restriction Measures

FR1(2)[3]_contracts	Categories	Definitions
1	Very Likely Financially Restricted	Information on Restructuring or Liquidation
2	Likely to be Financially Restricted	Only "Working Capital" and Average Interest Rate > 70% (80%) [90%] percentile and average maturity lower than 30% (20%) [10%] percentile
3	Not Clear	Not sufficient information to classify
4	Unlikely to be Financially Restricted	Financing and Average Interest Rate < 30% (20%) [10%] percentile and average maturity higher than 70% (80%) [90%] percentile
5	Very Unlikely to be Financially Restricted	Investment or Project Financing

## Panel B: Extensions of Core Financial Restriction Measures: From not enough information (3) to likely (2) to be financially restricted

3→2	Name
Loans 90 days delinquency	FR1(2)[3]_contracts_delinquency
>70% of portfolio of loans classified as bad loans	FR1(2)[3]_contracts_bad_portfolio
assets < 30% percentile and interest coverage < 30% percentile and fixed assets/Assets < 30% percentile	FR1(2)[3]_contracts_balancesheet
number of financial institutions relationships < 30% percentile	FR1(2)[3]_contracts_qifs
average colateral/assets < 30% percentile	FR1(2)[3]_contracts_collateral

*Notes:* All our measures are calculated every year, so they can change in our sample period. Our core measures presented in Table 3 Panel A are FR1\_Contracts, FR2\_Contracts, FR3\_Contracts. We build other financial restriction measures that complement these measures using information of firms related to the number of delinquency days of their loans; the demand of foreign exchange derivatives; number of bank relationships; information of portfolio of non-performing loans (bad loans); from average collateral provided; and from some balance sheet information. Panels B and C show financial restriction measures that complement those in Panel A.

Table 3: Financial Restriction Measures and Categories of Financial Restrictions

Panel C: Extensions of Core Financial Restriction Measures: From not enough information(3) to unlikely(4) to be financially restricted

3->4	Name
Outstanding Derivatives Contracts	FR1(2)[3]_contracts_derivatives
<30% of portfolio of loans classified as bad loans	FR1(2)[3]_contracts_portfolio
assets>30% percentile and interest coverage>30% percentile and fixed assets/Assets>30% percentile	FR1(2)[3]_contracts_balancesheet
number of financial institutions relationships>70% percentile average	FR1(2)[3]_contracts_qifs
colateral/assets>70% percentile	FR1(2)[3]_contracts_collateral

*Notes:* All our measures are calculated every year, so they can change in our sample period. Our core measures presented in Table 3 Panel A are FR1\_Contracts, FR2\_Contracts, FR3\_Contracts. We build other financial restriction measures that complement these measures using information of firms related to the number of delinquency days of their of loans; the demand of foreign exchange derivatives; number of bank relationships; information of portfolio of non-performing loans (bad loans); from average collateral provided; and from some balance sheet information. Panels B and C show financial restriction measures that complement those in Panel A.

Therefore, we have a total of 18 different measures of credit restriction of firms in our sample. To select only one that we think is more appropriate, we do the following. We estimate ordered probit panels, where the dependent variables are our measures of financial restrictions explained above and our explanatory variables are taken from [Whited & Wu index \(2006\)](#). They are the ratio of total debt to total assets, growth of firm's sales, the ratio of dividends to assets; and cash divided by total assets.<sup>13</sup>

We then select the best a priori classification of finance restriction as the one in which the average observed probabilities of forecasting all categories of financial restriction above, except for category 3, in our ordered panel probit estimations is the highest.

After selecting the best a priori classification, we define a firm as financially restricted in a certain year if its classification in that year is very likely (1) or likely (2) to be financially restricted. Otherwise, we consider a firm to be non-financially restricted if its classification falls in the categories likely (4) or very likely (5) to be financially non-restricted.

<sup>13</sup>We prefer [Whited & Wu \(2006\)](#) index from [Kaplan & Zingales \(1997\)](#) index because the former does not use Q of Tobin, which is important in our empirical analyses, because we have only private firms in our database.

## 4.2 Investment Demand

We use our selected financially restricted measures to estimate investment demand regressions of our unbalanced panel of firms from 2012 to 2020. We have different estimations depending on if the group is made of financially restricted or non-restricted firms.

In our main empirical analyses, we estimate Equation 1, using panel data and firm fixed effects. Equation 1 is based on [Gala et al. \(2020\)](#). [Gala et al.](#) estimate investment policy functions under general assumptions about technology and markets. They show that their policy functions are easy to estimate and, in addition, summarize the key predictions of any dynamic investment model. Because their method does not rely on Tobin’s Q, it does not require information about market values and can be applied to analyze private firms.

The dependent variable is the ratio between capex and assets. We estimate this equation for two distinct samples of firms and two distinct sample periods. In the case of the former, we estimate for financially restricted firms (FR) and non-financially restricted firms (NFR). In the case of the latter, in some estimations, we use the full sample period from 2012 to 2020, and in other ones, we implement difference-in-difference (hereafter dif-in-dif) estimations related to natural exogenous shock of the Covid-19 pandemic and the sample period goes from 2018 to 2020. In all estimations, we correct for heteroscedasticity.

In the case of estimations of the whole sample period, in addition to a constant, our regressors are the growth of operational revenue (`var_oper_rev`), the natural logarithm of total assets (`lassets`), and fixed total assets (`fixed_assets`).<sup>14</sup> In the case of dif-in-dif estimations, we include—in addition to the regressors of the estimations of the whole sample period—a dummy regressor, `pandemic`, equal to one in 2020 and zero in 2018 and 2019, and a regressor that is an interaction between `pandemic` and growth of operational revenue, `pandemic*var_oper_rev`. With the inclusion of this regressor, we want to verify if the BCB credit policies in 2020 were effective for firms that we classify as financially restricted. Table 4 presents a list of some of the more important BCB credit policies for firms in 2020.

The null hypotheses that we want to test are—following [Fazzari et al. \(1988\)](#)—that for the group of financially restricted firms, the coefficient ( $\beta_1$ ) of `var_oper_rev` is positive, while for the group of non-financially restricted firms either this coefficient is zero or negative. Furthermore, we want to attest if the credit policies of BCB during the pandemic decreased the sensitivity of investment to cash flow, that is we want to test if the coefficient ( $\beta_4$ ) of `pandemic*var_oper_rev` is negative.

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<sup>14</sup>Firm fixed effects are represented by  $a_i$ . Due to space restrictions, we do not show the estimated coefficients of constants.

Table 4: Some Important Credit Policies of Banco Central do Brasil during Covid-19 Pandemic (2020)

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Working capital programs to preserve business continuity (CGPE)
Purchase of private securities by BCB in the secondary market
Deduction on reserve requirement on savings deposits conditional on credit provision to micro and small companies
Real estate may be used as collateral in more than one credit operation
Emergency program provides payroll financing to SME in order to preserve employment in the segment
Fostering credit for small and medium-size enterprises
Relaxed provisions rules for refinancing loans of SME for six months

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$$\frac{\text{capex}_{i,t}}{\text{assets}_{i,t-1}} = \beta_0 + \beta_1 \text{var\_oper\_rev}_{i,t} + \beta_2 \text{lassets}_{i,t} + \beta_3 \text{fixed\_assets}_{i,t} + \beta_4 \text{pandemic}_t * \text{var\_oper\_rev}_{i,t} + \beta_5 \text{pandemic}_t + a_i + u_{i,t} \quad (1)$$

In the next section, we will present the results of our main empirical analyses and robustness exercises.

## 5 Results

### 5.1 Financial Restrictions Measures

Table 5 shows the average probabilities of the best model's forecasts of the categories of our ordered dependent variables. The FR3\_contracts\_qifs is the financial restriction measure with the highest probability of fitting (0.43), followed by FR2\_qifs (0.39).

Table 5: Selection of Financial Restriction Measures

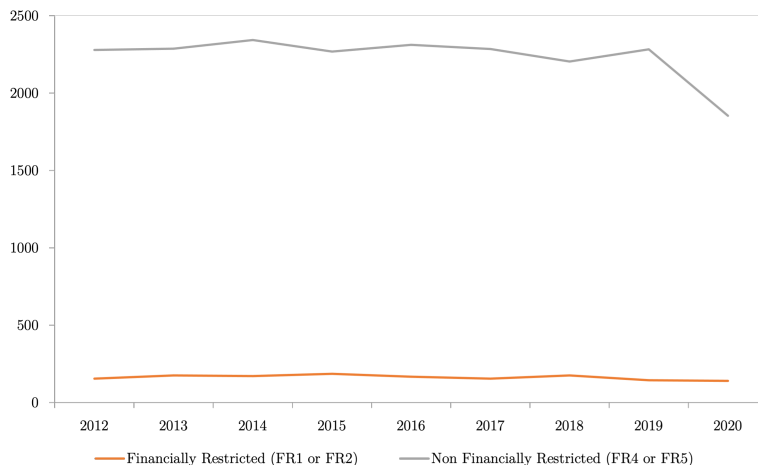
FR	Average Prob (FR=1 or 2 or 4 or 5)
Fr3_contracts_qifs	0.43
Fr2_contracts_qifs	0.39
Fr1_contracts_qifs	0.38

*Notes:* We estimate ordered probit panels, where the dependent variables are measures of financial restrictions and explanatory variables are taken from [Whited & Wu \(2006\)](#). The explanatory variables are ratio of total debt to total assets, growth of firm's sales, ratio of dividends to assets; and cash divided by total assets.

We define a firm as financially constrained (FR) firm if it belongs to FR3\_contracts\_qifs categories of very likely (1) or likely (2) to be credit constrained. We define a non-financially constrained firm (NFR) if it belongs to category likely (4) or very likely (5) to be non-financially constrained. Figure 4 displays the evolution in our sample period

of the number of private firms in each category of financial restriction measures based on FR3\_contracts\_qifs.<sup>15</sup> One can see that categories financially restricted (FR) non-financially restricted (NFR) are relatively stable in our sample period.

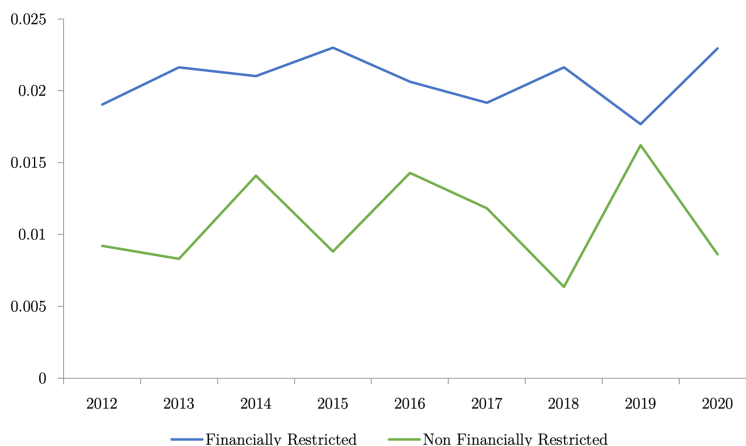
Figure 4: Numbers of Financial Restricted and Non-Financially Restricted Firms



Source: Banco Central do Brasil (BCB) – System of Credit Information (SCR)

Figure 5 shows the year average classification of financially constrained and non-financially firms in our sample period. Both series look very stable throughout the years. Figure 6 and Figure 7 show the number of financially constrained firms and non-financially restricted firms respectively classified by sectors. It is clear from both Figures, that firms of the services sector are the ones with more firms in both categories.

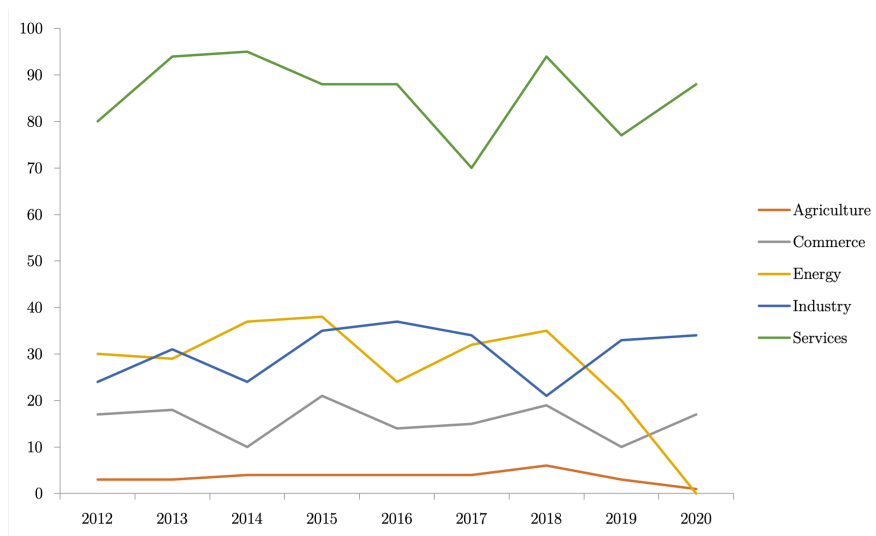
Figure 5: Average Financially Restricted (FR) and Non-Financially Restricted (NFR) Firms



Source: Banco Central do Brasil (BCB) – System of Credit Information (SCR)

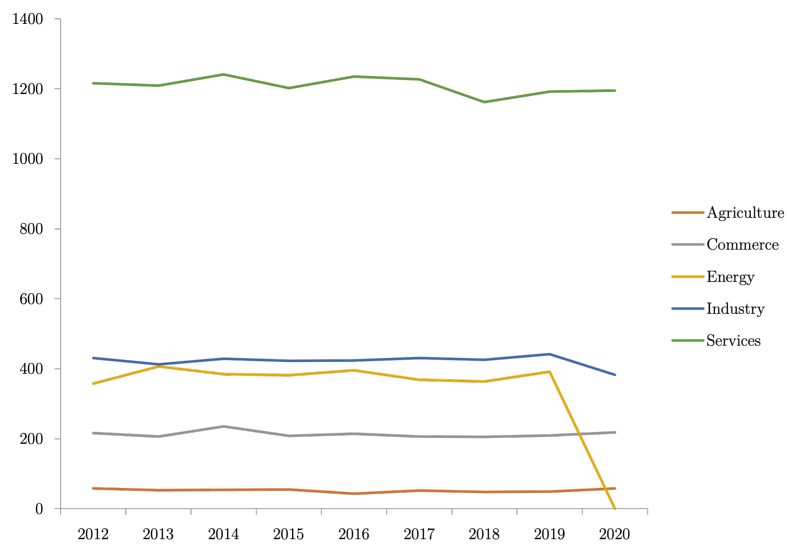
<sup>15</sup>FR(1 to 5) means that FR3\_Contracts\_qifs is equal to 1 to 5 respectively.

Figure 6: Number of Financial Restricted Firms and Sectors of the Economy



Source: Banco Central do Brasil (BCB) – System of Credit Information (SCR)

Figure 7: Number of Non-Financially Restricted Firms and Sectors of the Economy



Source: Banco Central do Brasil (BCB) – System of Credit Information (SCR)

We perform mean tests of some characteristics of group of firms financially restricted or not. Table 6 shows that the group of firms that are non-financially restricted are larger and invest more than firms that are financially restricted. However, we do not see any statistically significant difference between the groups of financially restricted or not as far as profitability (ebitda/assets) and growth of operational reserves are concerned (var\_oper\_rev).

Table 6: Mean Tests of Difference in Characteristics of Financially Restricted and Non-Financially Restricted

	(FR=1-FR=0)
logassets	-0.2119***
capex/assets	-0.5500***
ebitda/assets	1.3883
var_oper/assets	0.2385
var_oper_rev	0.0501
leverage	0.9757***

## 5.2 Results of Estimations of Investment Functions

Table 7 displays the results of our main estimations of investment demand, Equation 1, based on Gala et al. (2020) for financially restricted and non-financially restricted firms, respectively. The results in Table 7 show that the sensitivity of investment in relation to cash flow (measured by the coefficient of growth of operational revenue(var\_oper\_rev)) for financially restricted firms is positive and statistically significant and varies from a minimum of 0.0048 to a maximum of 0.013, capex\_assets. Therefore a 1% increase in cash flow ceteris paribus would increase from a minimum of 0.4% to a maximum of 1.38% the ratio of capex to assets. The coefficients estimated of the var\_oper\_rev for the case non-financially restricted firms is negative and statistically significant, -0.211, or non-statistically significant.

In addition, the dif-in-dif estimated regressions indicate that, in the pandemic period, the coefficient of the interaction between cash flow measure and pandemic, var\_oper\_rev\* pandemic, is not statistically significant. This may point to the fact that credit policies of BCB, during the pandemic, directed towards small and medium size Brazilian firms, such as the ones we have in our sample, did not mitigate the effects of credit restrictions on investment of these firms.

In the next section, we will present results of some robustness exercises to verify if the results of our main empirical analyses described above continue valid.

Table 7: Main Investment Demand Estimations based on [Gala et al. \(2020\)](#)

	(Capex/Assets)			
	FR	NFR	FR	NFR
var_oper_rev	0.0048*** (3.4994)	-0.211*** (-13.46)	0.0138 (2.1207)	-0.000 (-0.407)
lassets	0.0074 (0.7242)	-0.908** (-2.060)	0.0016 (0.1619)	-0.006 (-1.979)
imob_assets	0.1012 (0.9559)	-0.608 (-1.390)	0.0431 (1.5766)	0.0742 (4.0823)
pandemic*var_oper_rev			60.0019 (0.1438)	-0.000 (-1.358)
pandemic			-0.036 (-2.807)	-0.006 (-0.925)
Robust Covariance	yes	yes	yes	no
Fixed Effects	yes	yes	yes	yes
Sample Period	2012-2020	2012-2020	2018-2020	2018-2020
Dif-in-Dif	no	no	yes	yes
N	213	1435	84	651
R2	0.029	0.93	0.16	0.17

t-statistics under parentheses

\*p&lt;0.10 \*\* p&lt;0.05 \*\*\* p&lt;0.001

*Notes:* We estimate Equation 1 in the text, using panel data and firm fixed effects. Equation 1 is based on [Gala et al. \(2020\)](#). [Gala et al.](#) estimate investment policy functions under general assumptions about technology and markets and do not use average Q of Tobin. FR means financial restrictions, while NFR means no financial restrictions.

### 5.3 Robustness Exercises of the Estimation of Investment

The robustness exercises comprise estimations of different specifications in line with the econometric model of [Gala et al. \(2020\)](#); estimation of different models of investment based on [Bond et al. \(2003\)](#) and [Eberly et al. \(2012\)](#); estimations dealing with possible endogeneity of growth of operational revenues and measurement errors of financial restrictions measures.

#### 5.3.1 Robustness Exercises based on [Gala et al. \(2020\)](#)

We estimate several regressions like Equation 1 above and centered on [Gala et al. \(2020\)](#), correcting for outliers of investment, substituting some regressors by other ones, changing the definitions of regressions and including new ones.



In Table 8, Panels A to F, we show the results of these robustness exercises: in Panel A, we restrict our sample of firms to those that have ratio of capex to assets between the 5% percentile and 95% of the cross-section distribution of this ratio every year in our sample period; in Panel B, we substitute growth of operational revenue for growth of ebitda; in Panel C, we include macroeconomic regressors, such as SELIC rate, GDP growth and Inflation<sup>16</sup>; in Panel D, we include dummy regressors indicating sectors of the economy (agriculture, commerce, energy, industry and services)<sup>17</sup>; in Panel E, we include regressors leverage and liquidity (measured by the ratio of cash to assets); in Panel F, we estimate Equation 1 with the same regressors using deflated (real) values, instead of nominal ones.<sup>18</sup>

The results displayed in Table 8 Panels A to F confirm those we obtain in our main empirical analyses, that is, the elasticity of investment in relation to cash flow for financially restricted firms is positive and statistically significant, and negative or statistically non-significant for the case non-financially restricted firms. In addition, in most dif-in-dif estimated regressions, the coefficient of the interaction between cash flow measure and pandemic, is not statistically significant.

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<sup>16</sup>Inflation is measured by “Índice de Preços ao Consumidor Ampliado” (hereafter IPCA).

<sup>17</sup>In the case of these regressions, we use between effects instead of fixed effects.

<sup>18</sup>To deflate nominal values we use price index IPCA.

Table 8: Robustness Exercises of Investment Demand Estimations based on  
[Gala et al. \(2020\)](#)

Panel A: Excluding Outliers of Capex/Asset

	(Capex/Assets)			
	FR	NFR	FR	NFR
var_oper_rev	0.0048*** (3.4990)	-0.000* (-1.817)	0.0125* (1.8985)	6.0617 (0.2511)
lassets	0.0074 (0.7241)	-0.000 (-0.031)	-0.000 (-0.069)	-0.002 (-1.140)
imob_assets	0.1012 (0.9558)	-0.126*** (-5.930)	0.0291 (1.2563)	0.0415*** (4.1004)
pandemic*var_oper_rev			0.0019 (0.1472)	-0.000* (-1.853)
pandemic			-0.030*** (-2.685)	-0.006 (-1.493)
Robust Covariance	yes	yes	yes	yes
Fixed Effects	yes	yes	yes	yes
Sample Period	2012-2020	2012-2020	2018-2020	2018-2020
Dif-in-Dif	no	no	yes	yes
N	210	1417	83	639
R2	0.029	0.023	0.16	0.056

t-statistics under parentheses

\*p<0.10 \*\* p<0.05 \*\*\* p<0.001

*Notes:* We estimate several regressions like Equation 1 above and centered on [Gala et al. \(2020\)](#), correcting for outliers of investment, substituting some regressors by other ones, changing their definitions and including new regressors. FR means financial restrictions, while NFR means no financial restrictions.

Table 8: Robustness Exercises of Investment Demand Estimations based on [Gala et al. \(2020\)](#)

Panel B: Substituting Growth of Operational Revenues for Growth of Ebitda

	(Capex/Assets)			
	FR	NFR	FR	NFR
var_ebitda	92.455*** (4.2789)	0.0002 (0.9823)	-0.000 (-0.059)	-3.536 (-0.096)
lassets	0.0430** (2.4855)	-5.188 (-1.461)	0.0068 (0.7111)	-0.008** (-2.450)
imob_assets	0.0907 (0.8991)	1.8470*** (3.0275)	0.0319 (1.1043)	0.0728*** (3.9640)
pandemic*var_oper_rev			0.0163 (1.2172)	-0.000* (-1.867)
pandemic			-0.036 (-2.543)	-0.007 (-0.959)
Robust Covariance	yes	yes	yes	yes
Fixed Effects	yes	yes	yes	yes
Sample Period	2012-2020	2012-2020	2018-2020	2018-2020
Dif-in-Dif	no	no	yes	yes
N	210	1417	80	602
R2	0.029	0.023	0.67	0.52

t-statistics under parentheses

\*p<0.10 \*\* p<0.05 \*\*\* p<0.001

*Notes:* We estimate several regressions like Equation 1 above and centered on [Gala et al. \(2020\)](#), correcting for outliers of investment, substituting some regressors by other ones, changing their definitions and including new regressors. FR means financial restrictions, while NFR means no financial restrictions.

Table 8: Robustness Exercises of Investment Demand Estimations based on [Gala et al. \(2020\)](#)

Panel C: Including Macroeconomic Regressors: SELIC, GDP and Inflation

	(Capex/Assets)			
	FR	NFR	FR	NFR
var_oper_rev	0.0022*	-0.198***	0.0123*	-4.498
	(1.6415)	(-12.94)	(2.4944)	(-0.014)
lassets	-0.002	-0.335***	0.0094	0.0005
	(-0.302)	(-8.963)	(1.0660)	(0.1720)
imob_assets	0.0597	0.2329	0.0631**	0.091***
	(0.8705)	(0.7190)	(2.5548)	(4.5999)
pandemic*var_oper_rev			0.0067	-0.000
			(0.5356)	(-1.314)
pandemic			-0.034***	-0.016
			(-2.765)	(-2.073)
Macroeconomic variables: GDP, Inflation and SELIC	yes	yes	yes	yes
Robust Covariance	yes	yes	yes	yes
Fixed Effects	yes	yes	yes	yes
Sample Period	2012-2020	2012-2020	2018-2020	2018-2020
Dif-in-Dif	no	no	yes	yes
N	198	1318	84	651
R2	0.067	0.078	0.47	0.12

t-statistics under parentheses

\*p<0.10 \*\* p<0.05 \*\*\* p<0.001

*Notes:* We estimate several regressions like Equation 1 above and centered on [Gala et al. \(2020\)](#), correcting for outliers of investment, substituting some regressors by other ones, changing their definitions and including new regressors. FR means financial restrictions, while NFR means no financial restrictions.

Table 8: Robustness Exercises of Investment Demand Estimations based on [Gala et al. \(2020\)](#)

Panel D: Including Regressors Indicating Sectors

	(Capex/Assets)			
	FR	NFR	FR	NFR
var_oper_rev	0.0048 (0.8592)	-0.211** (-2.138)	-4.498 (-0.016)	-4.498 (-0.010)
lassets	0.0074 (0.4045)	-0.908* (-1.742)	0.0005 (0.1870)	0.0005 (0.1844)
imob_assets	0.1012 (0.4807)	-0.608 (-1.423)	0.0912*** (4.4628)	0.0912 (4.5473)
pandemic*var_oper_rev			-0.000 (-0.517)	-0.000 (-0.510)
pandemic			-0.016** (-2.176)	-0.016 (-2.016)
Sectors	yes	yes	yes	yes
Robust Covariance	yes	yes	yes	yes
Fixed Effects	yes	yes	yes	yes
Sample Period	2012-2020	2012-2020	2018-2020	2018-2020
Dif-in-Dif	no	no	yes	yes
N	213	1435	84	651
R2	0.029	0.93	0.22	0.08

t-statistics under parentheses

\*p<0.10 \*\* p<0.05 \*\*\* p<0.001

*Notes:* We estimate several regressions like Equation 1 above and centered on [Gala et al. \(2020\)](#), correcting for outliers of investment, substituting some regressors by other ones, changing their definitions and including new regressors. FR means financial restrictions, while NFR means no financial restrictions.

Table 8: Robustness Exercises of Investment Demand Estimations based on [Gala et al. \(2020\)](#)

Panel E: Including Regressors Indicating Sectors

	(Capex/Assets)	
	FR	NFR
var_oper_rev	0.0054*** (4.1072)	-0.211*** (-13.45)
lassets	0.0063 (0.6224)	-0.907** (-2.058)
imob_assets	0.1017 (0.9556)	-0.617 (-1.412)
Robust Covariance	yes	yes
Fixed Effects	yes	yes
Sample Period	2012-2020	2012-2020
Dif-in-Dif	no	no
N	213	1435
R2	0.03	0.93

t-statistics under parentheses

\*p<0.10 \*\* p<0.05 \*\*\* p<0.001

*Notes:* We estimate several regressions like Equation 1 above and centered on [Gala et al. \(2020\)](#), correcting for outliers of investment, substituting some regressors by other ones, changing their definitions and including new regressors. FR means financial restrictions, while NFR means no financial restrictions.

Table 8: Robustness Exercises of Investment Demand Estimations based on [Gala et al. \(2020\)](#)

Panel F: Including Regressors Leverage and Liquidity

	(Capex/Assets)			
	FR	NFR	FR	NFR
var_oper_rev	0.0033** (2.1138)	-0.220*** (-30.08)	0.0093 (1.6032)	-8.716 (-0.347)
lassets	0.0130 (1.1093)	-0.167 (-0.973)	0.0018 (0.1780)	-0.007 (-2.106)
imob_assets	0.1102 (1.2590)	-0.475 (-0.918)	0.0273 (0.9618)	0.0753*** (4.0570)
pandemic*var_oper_rev			-0.008 (-1.477)	-0.000 (-0.850)
pandemic			-0.038 (-3.054)	-0.005 (-0.738)
Robust Covariance	yes	yes	yes	yes
Fixed Effects	yes	yes	yes	yes
Leverage and Liquidity	yes	yes	yes	yes
Sample Period	2012-2020	2012-2020	2018-2020	2018-2020
Dif-in-Dif	no	no	yes	yes
N	207	1407	80	630
R2	0.065	0.96	0.10	0.07

t-statistics under parentheses

\*p<0.10 \*\* p<0.05 \*\*\* p<0.001

*Notes:* We estimate several regressions like Equation 1 above and centered on [Gala et al. \(2020\)](#), correcting for outliers of investment, substituting some regressors by other ones, changing their definitions and including new regressors. FR means financial restrictions, while NFR means no financial restrictions.

### 5.3.2 Different Investment Specifications: Vector Error Correction and Lagged Investment

We estimate, using [Arellano & Bond \(1991\)](#), investment demand functions taking in consideration lagged investment effect (one lag of investment as one of the regressors) based on [Eberly et al. \(2012\)](#). The authors show that the best predictor of current investment at the firm level is lagged investment. In their view, lagged-investment effect is empirically more important than the cash flow and Q effects combined.

We also estimate a vector error correction model of investment (see [Bond et al. \(2003\)](#)), where there is an error correction vector of operational revenues and fixed assets lagged two periods. [Bond et al.](#) construct company panel data sets for manufacturing firms

in Belgium, France, Germany, and the United Kingdom, covering the period 1978-1989. They use these data to estimate empirical investment equations, and to investigate the role played by financial factors in each country.

The results presented in Table 9 for Eberly et al. (2012) and Table 10 for Bond et al. (2003) models, once more, endorse the ones of our main empirical analyses.<sup>19</sup>

Table 9: Investment Demand Estimations: Lagged Effects Based on Eberly et al. (2012)

	(Capex/Assets)			
	FR	NFR	FR	NFR
capex_assets(-1)	-0.087 (-0.27)	-0.005 (-0.688)	0.3525* (1.8040)	-0.029** (-2.442)
var_oper_rev	0.0003 (0.2575)	-0.217*** (-15.01)	0.0162*** (3.5393)	-0.00001 (-0.334)
lativo	-0.024** (-2.090)	0.1991 (0.4309)	0.0068 (0.5865)	-0.009* (-1.676)
imob_assets	-0.440 (-1.586)	-0.384 (-0.444)	0.0501 (1.5258)	0.0711*** (3.6965)
pandemic*var_rec			0.0240 (1.1936)	-0.0001 (-0.482)
pandemic			-0.057*** (-2.886)	-0.029** (-2.442)
Robust Covariance	yes	yes	yes	no
Arellano-Bond	yes	yes	yes	yes
Fixed Effects	yes	yes	yes	yes
Sample Period	2012-2020	2012-2020	2018-2020	2018-2020
Dif-in-Dif	no	no	yes	yes
J Statistic	19.79	28.80	23.80	24.79
N	213	1435	52	407
R2	0.029	0.93	0.92	0.00012

t-statistics under parentheses

\*p<0.10 \*\* p<0.05 \*\*\* p<0.001

*Notes:* We estimate, using Arellano & Bond (1991), investment demand functions taking in consideration lagged investment effect (one lag of investment as a regressor) based on Eberly et al. (2012). FR means financial restrictions, while NFR means no financial restrictions.

<sup>19</sup>In the case of the Bond et al. (2003) model, the coefficient of var\_rev\_oper is not statistically significant for firms with financial restrictions and, in accordance with our null hypothesis, negative and statistically significant for firms with no financial restrictions



Table 10: Investment Demand Estimations based on [Bond et al. \(2003\)](#)

	(Capex/Assets)	
	FR	NFR
capex_assets(-1)	-0.745 (-1.629)	-0.006 (-1.173)
var_oper_rec	-0.000 (0.174)	-0.213*** (-17.89)
lassets	-0.010 (-0.304)	-0.006 (1.6130)
imob_assets	0.010 (0.35)	-0.983*** (-2.92)
var_oper-fixed(-2)	-0.010 (-0.16)	7.5E-10** (4.93)
Robust Covariance	yes	yes
Arellano-Bond	yes	yes
Sample Period	2012-2020	2012-2020
Fixed Effects	yes	yes
N	86	666
J Statistic	38.6	26.04

t-statistics under parentheses

\*p<0.10 \*\* p<0.05 \*\*\* p<0.001

*Notes:* We estimate a vector error correction model of investment based on [Bond et al. \(2003\)](#), where there is an error correction vector of sales and fixed assets lagged two periods, oper\_rev-fixex(-2) in addition to the regressors of our main specification Equation 1. FR means financial restrictions, while NFR means no financial restrictions.

### 5.3.3 Endogeneity

In a fourth attempt to verify the robustness of our main empirical results, we contemplate possible endogeneity of the regressor growth of operational revenues in the estimation of Equation 1. We perform two different econometric exercises to deal with this possibility. In the first one, we estimate Equation 1 using two stages least squares (hereafter 2SLS) with the first lag of growth of operational revenue as an instrument.<sup>20</sup> The results shown in Table 11 Panel A confirm once more our main empirical results

In the second econometric exercise, we implement Average Treatment Effects (hereafter ATE) estimations based on Nearest Neighbor Matching (hereafter NN) and Propensity Matching Score (hereafter PMS). In the case of NN, the treatment variable is a

<sup>20</sup>We test other instruments. However, [Hausman's 1978](#) test of specification rejects all estimations with more lags in favor of regression with one lag of growth of growth of operational revenue as an instrument.

dummy variable that indicates a firm is financially restricted (FR), matching variable is natural logarithm of assets and the regression is identical to Equation 1. In the case of PMS, the matching is based on a regression of a dummy that indicates financial restriction, FR, with respect to the same regressors we use in the selection of the best financial restriction measure based on [Whited & Wu \(2006\)](#)<sup>21</sup> and the regression specification is the same as Equation 1.

The results of our estimations—which deal with possible endogeneity of growth of operational revenues—are presented in Table 11 Panels A and B. As one can observe, the coefficients estimated in 2SLQ estimations confirm our main empirical analyses results. In the case of the ATE estimations, one can see that the ATE of the ratio of capex to assets is negative and statistically significant in NN estimation for the pandemic period. In other estimations, ATE sign is also negative but not statistically significant.

Table 11: Investment Demand Estimations: Taking in Consideration Endogeneity

Panel A: Lag of growth of operational revenue as instrument

	(Capex/Assets)			
	FR	NFR	FR	NFR
var_oper_rev	0.0048*** (3.4994)	-0.211*** (-13.46)	0.0200*** (3.4662)	-0.000 (-0.604)
lativo	0.0074 (0.7242)	-0.908** (-2.060)	0.0055 (0.0142)	-0.007 (-1.429)
imob_assets	0.1012 (0.9559)	-0.608 (-1.390)	0.0886** (2.2838)	0.0927*** (5.1567)
pandemic*var_oper_rev			-0.000767 (-0.043)	-0.00027 (-0.230)
pandemic			-0.06603*** (-2.915)	-0.02869*** (-2.366)
pandemic*var_oper_rev+var_oper_rev			0.0019 (1.1659)	-0.0009 (-1.48)
Robust Covariance (cluster sectors)	yes	yes	yes	no
Instruments	var_rev_oper(-1)	var_rev_oper(-1)	var_rev_oper(-1)	var_rev_oper(-1)
Fixed Effects	yes	yes	yes	yes
Sample Period	2012-2020	2012-2020	2018-2020	2018-2020
Dif-in-Dif	no	no	yes	yes
N	213	1435	52	407
R2	0.029	0.93	0.92	0.0012
F	8.71	5935.47	4.57	7.58

t-statistics under parentheses

\*p<0.10 \*\* p<0.05 \*\*\* p<0.001

*Notes:* In panel A, we show the estimation of Equation 1 using 2SLS with the first lag of growth of operational revenue as an instrument. In panel B, we display the results of Average Treatment Effects estimations based on Nearest Neighbor Matching (NN) and Propensity Matching Score (PMS). In the case of NN, the treatment variable is a dummy variable that indicates a firm is financially restricted (FR), the matching variable is natural logarithm of assets, and the regression is identical to Equation 1. In the case of PMS, the matching is based on a regression of a dummy that indicates financial restriction, FR, with respect to the same regressors we use in the selection of the best financial restriction measure based on [Whited & Wu \(2006\)](#) and the regression specification is the same as Equation 1.

<sup>21</sup>See the data section of this paper.

Panel B: Average Treatment Effects

	ATE (Capex/Assets)	
	FR=1-FR=0	(FR=1-FR=0)
Nearest Neighbour Matching	-0.13 (-1.03)	-0.016* (-1.71)
Propensity Matching Score	-0.14 (-1.06)	-0.0079 (-0.81)
Pandemic	No	Yes
OBS	1544	188
Number of Matches	2	2

t-statistics under parentheses

\*p<0.10 \*\* p<0.05 \*\*\* p<0.001

*Notes:* In panel A, we show the estimation of Equation 1 using 2SLS with the first lag of growth of operational revenue as an instrument. In panel B, we display the results of Average Treatment Effects estimations based on Nearest Neighbor Matching (NN) and Propensity Matching Score (PMS). In the case of NN, the treatment variable is a dummy variable that indicates a firm is financially restricted (FR), the matching variable is natural logarithm of assets, and the regression is identical to Equation 1. In the case of PMS, the matching is based on a regression of a dummy that indicates financial restriction, FR, with respect to the same regressors we use in the selection of the best financial restriction measure based on [Whited & Wu \(2006\)](#) and the regression specification is the same as Equation 1.

### 5.3.4 Errors of Financial Restriction Measures

Another reasonable and relevant concern of our main results presented in Table 6 is related to our identification of financially restricted firms. Our criteria can be questioned in several manners. It can be wrong, or it may not describe all the possibilities in which firms could be subject to financial constraints.

To handle possible measurement errors, we implement two distinct set of exercises. In the first set, we analyze in detail the group of firms that we were unable to classify into more likely or less likely to be financially constraint due to lack of information, that is, firms in category 3 of our chosen financially restricted measure. In the second set of exercises, we estimate semi-parametric models based on [Klein & Spady \(1993\)](#), where the dependent variables are dummies indicating financial restrictions (equal to one) or not (equal to zero) of firms, depending on our measures of financial restrictions presented earlier in the paper.<sup>22</sup>

<sup>22</sup>[Hausman et al. \(1998\)](#) show that misclassification of dependent variables in a discrete-response model causes inconsistent coefficient estimates when traditional estimation techniques, such as probit or logit, are used. One of their suggestions to solve this problem is to use semi-parametric models, such as [Klein & Spady \(1993\)](#).

### 5.3.4.1 Sample of Firms with Not Enough Information to Know its Credit Status

Firms in category 3 of `FR3_contracts_qifs` (our chosen measure of financial restrictions) are the ones in which the data that we have does not allow us to observe some characteristics that would lead us to classify them in either the category likely to be financially constrained or likely to be non-financially constrained. Firms in these situations could be the ones where there are not investment opportunities available, or those that our database of loan contracts from SCR are unable to give us an understanding of how they finance their investment.

Figure 8 describes the number of firms that we cannot discriminate into some degree of financially restricted or not. The services sector is by far the sector with more firms of this kind in our sample. It is interesting to note that the number of firms in this category increases overtime, which does not happen to firms in other sectors, whose numbers show a stable dynamic or a decreasing trend. In the pandemic period, in 2020, there is a clear increase in the number of firms of service sector in category 3. Considering that the services sector in Brazil (and worldwide) was the one that suffered the most with the pandemic in terms of growth, to us this seems to be a first and important evidence towards considering firms in category 3 as those with very few investment opportunities.

In addition, Figure 9 shows the percentage of firms in category 3 that have capex higher than zero in our sample period. As one can easily see, very few firms in this category invest in our sample period, which is further indication these firms may be experiencing few incentives to invest.

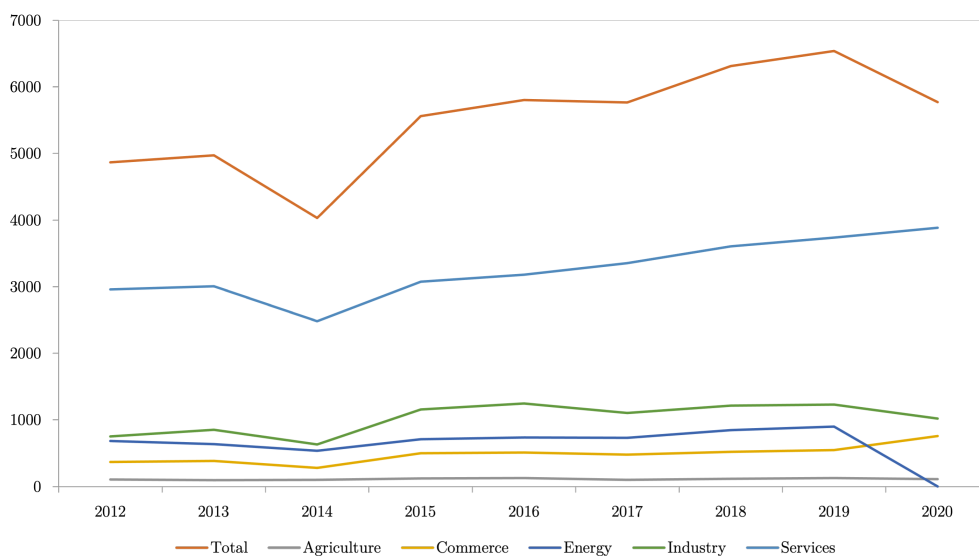
Figure 10 corroborates, in our view, our understanding—based on Figures 8 and 9—of few investment opportunities for firms in category 3. It shows the number of firms in this category separated in the following manner: those that have capex greater than zero only; those that have operational revenues greater than zero only; those that have cash and operational revenues greater than zero at the same time—but not capex greater than zero—and, finally, those that have capex and operational revenues greater than zero at the same time—but not cash greater than zero.

Two observations are striking in Figure 10: there are very few firms that have capex greater than zero only relative to ones that have operational revenues only greater than zero; there are very few firms that have capex and operational revenues greater than zero at same time in relation to those that have operational revenues and cash greater than zero at the same time. The two observations put together are, in our opinion, *prima facie* evidence that the firms in our sample are hoarding cash instead of investing, and this is probably due to lack of good projects to invest in.

Therefore, evidence presented in Figures 8 to 10 lead us to conclude that we are not able to identify firms in category 3 of our chosen financial restriction measure because these firms are not, probably, demanding investment, therefore not needing credit from

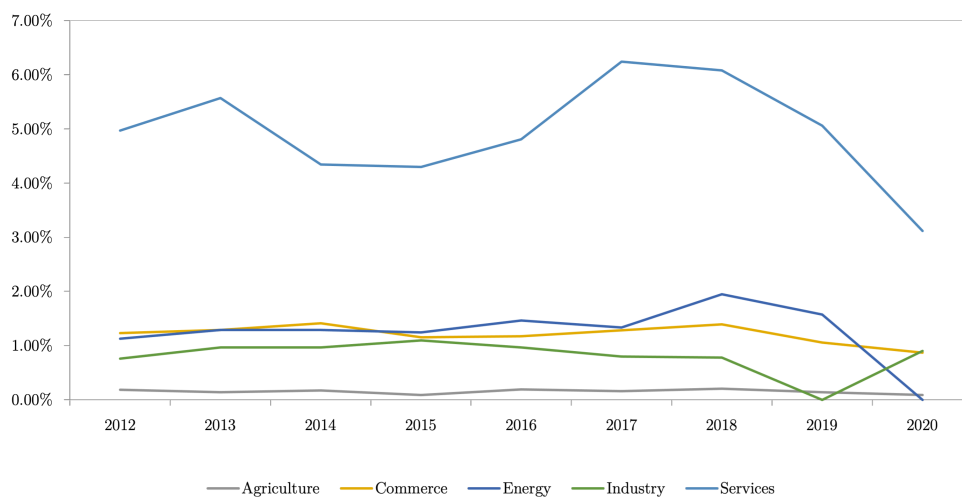
financial institutions for this purpose.

Figure 8: Number of Firms in Category 3



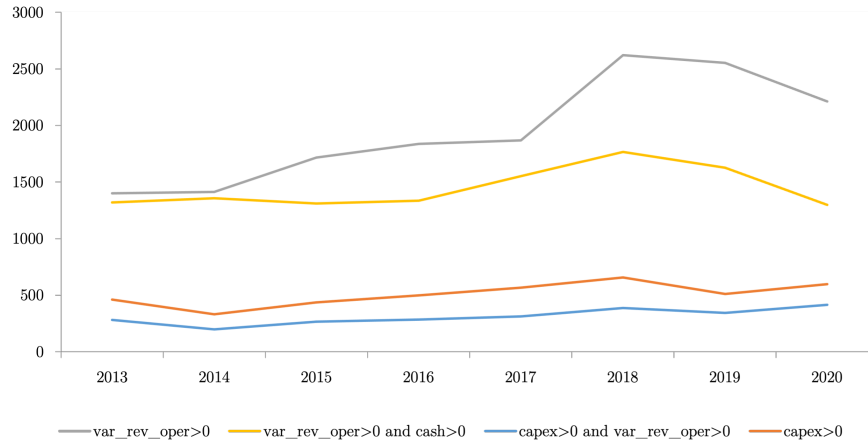
Source: Banco Central do Brasil (BCB) – System of Credit Information (SCR)

Figure 9: Percentage of Firms in Category 3 with Capex Higher than Zero Separated by Sectors of the Economy



Source: Banco Central do Brasil (BCB) – System of Credit Information (SCR)

Figure 10: Number of Firms Considering Capex, Cash and Growth of Operational Revenues Greater Than Zero



Source: Banco Central do Brasil (BCB) – System of Credit Information (SCR)

To try to reinforce our conclusions mentioned above, we create a new measure of financially restricted firms for some firms that fall in category 3. This new measure is applied to those firms that have capex and operational revenues—but not cash—greater than zero at the same time. Thus, we are suggesting that they use operational revenues to invest, instead of asking for bank credit, that is, they are more likely to be financially constrained. Then we estimate equation 1 for the whole sample and for the pandemic period once more only for firms that fit in this situation. If there was evidence that these firms were financially constrained then we would observe a positive and statistically significant coefficient of growth of operational revenues, in line with Fazzari et al. (1988). As the results presented in Table 12 show coefficients estimated are negative and non-statistically significant, which points, again, to our understanding that available information on firms of category 3 do not make it possible to identify them in more or less likely to be financially restricted or not.

Table 12: Financial Restriction Estimation Including Firms in Category 3 with Capex and Growth of Operational Revenues Greater Than Zero

	(Capex/Assets)	
	FR	FR
var_oper_rev	-9.291 (-0.030)	-8.349 (-0.058)
lativo	-0.041 (-1.613)	0.0098 (1.5249)
imob_assets	-0.599*** (-1.730)	0.1598 (4.6589)
Robust Covariance (cluster sectors)	yes	yes
Fixed Effects	yes	yes
Sample Period	2012-2020	2018-2020
dif-in-dif	no	yes
N	489	219
R2	0.08	0.26

t-statistics under parentheses  
 \*p<0.10 \*\* p<0.05 \*\*\* p<0.001

*Notes:* We estimate Equation 1 in the text, using panel data and firm fixed effects only for firms in category 3 with capex and growth of operational revenue greater than zero. Equation 1 is based on [Gala et al. \(2020\)](#). [Gala et al.](#) estimate investment policy functions under general assumptions about technology and markets and do not use average Q of Tobin. FR means financial restrictions, while NFR means no financial restrictions.

### 5.3.4.2 Semi-parametric Estimation of [Klein & Spady \(1993\)](#)

The other set of exercises we perform to verify if there are error measurements in our definitions of financial restrictions is grounded on semi-parametric regressions based on [Klein & Spady \(1993\)](#). The authors implement an estimator for a semi-parametric binary choice model. The distribution function of error of the regression is unknown and estimated nonparametrically by kernel or local linear regression using the quartic kernel. The [Klein & Spady](#) estimator can only identify coefficients up to location and scale, so the one of the coefficients is normalized to 1. The bandwidth is computed by [Silverman's \(1986\)](#) rule.

To perform [Klein & Spady](#) regressions with our financial measures, we define correspondent measures of financial restrictions of these measures – one if the firm is in liquidation or restructuring (category 1) or is very likely to be financially restricted (category 2) and zero in case the firms is unlikely (category 4) or very unlikely (category 5) to be financially constrained. We use the same regressors we use in our ordered probit regressions, shown in section 4. We are interested in measuring the mean squared errors (hereafter MSE) of the difference between observed and predicted dependent variables of the estimations for every measure of financial restrictions. Table 13 exhibits the values

of MSE and as one can easily see they are very small ones (all of them are lower than 10%), which is another evidence that our financial restrictions measures are less likely to have measurement errors.

Table 13: Considering Misclassification of Measures of Financial Restrictions: [Klein & Spady \(1993\)](#)

Measures of Financial Restrictions	Mean Squared Error
FR1_contracts	0.07068
FR2_contracts	0.00011
FR3_contracts	0.00129
FR1_contracts_derivatives	0.06887
FR2_contracts_derivatives	0.06966
FR3_contracts_derivatives	0.02212
FR1_contracts_portfolio	0.00993
FR2_contracts_portfolio	0.07453
FR3_contracts_portfolio	0.07204
FR1_contracts_balancesheet	0.07417
FR2_contracts_balancesheet	0.07250
FR3_contracts_balancesheet	0.03371
FR1_contracts_qifs	0.07218
FR2_contracts_qifs	0.09402
FR3_contracts_qifs	0.03525
FR1_contracts_collateral	0.05831
FR2_contracts_collateral	0.03109
FR3_contracts_collateral	0.03723

*Notes:* We estimate semi-parametric regressions based on [Klein & Spady \(1993\)](#). The authors implement an estimator for a semi-parametric binary choice model. The distribution function of error of the regression is unknown and estimated non parametrically by kernel or local linear regression using the Quartic kernel. The Klein and Spady estimator can only identify coefficients up to location and scale, so the one of the coefficients is normalized to one. The bandwidth is computed by [Silverman's \(1986\)](#) rule. We obtain mean squared errors (MSE) of the difference between observed and predicted dependent variables.

## 6 Conclusion

In this paper, we build original financial restriction measures. We use these measures to have an idea of financial restrictions of 8,071 firms in Brazil from 2012 to 2020. To build these measures, we use microdata of 3,469,135 loan contracts written between these firms and financial institutions in this period.

Our financial restrictions measures do a good job in explaining the credit market of some small and medium-size firms in Brazil and indicate that their investment is negatively related to the degree of their financial restrictions. Moreover, our results show that this did not change during the pandemic, despite the efforts of BCB to relax credit restrictions of these firms in this period.



We ponder that our paper contributes in relevant ways to empirical literature and in terms of policy. Considering the former, our paper creates original measures of financial restrictions that, we think, are more accurate than the ones that exist in the literature so far. As far as we know, our paper is the first one in the literature to construct these measures by analyzing information directly from loan contracts of firms with financial institutions. In terms of the latter, we ponder that our financial restriction measures can be used by policy makers to have a better understanding of the stance of the credit market for private firms in Brazil.

Future research could expand the number of private firms to be studied, which depends, of course, on the availability of detailed balance sheet and financial information of private joint stock firms and limited liabilities ones. Future research could also use firm's loans contracts in SCR to try to infer agency costs, due to information asymmetries and moral hazards, which, we cogitate, could improve the financial restriction measures that we build in this paper.

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